



DISTRIBUTION OF CORRELATED LOGNORMAL RANDOM SUM

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Abstract

Correlated lognormal random sum (CLRS) is random sum, $Z_N = X_1 + X_2 + \cdots + X_N$, $N = 0, 1, 2, \dots$, when X_1, X_2, \dots, X_N are correlated and the distribution of X_1, X_2, \dots, X_N given $N = l$ is multivariate lognormal. This paper discusses the approximation distribution for CLRS. The approximation percentile, mean and variance of CLRS are also discussed. Monte Carlo simulation is conducted to investigate the approximations. The simulation results show that the approximation distribution, percentile, mean and variance of CLRS are valid and good approximations.

1. Introduction

The sum of independent random variables also called the *random sum*, $Z_N = X_1 + X_2 + \cdots + X_N$, $N = 0, 1, 2, \dots$, where $Z_0 = 0$, has been

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discussed by many authors (for examples, see Taylor and Karlin [12], and Klugman et al. [6]). It is assumed that X_1, X_2, \dots, X_N are independent random variables, and N is a discrete random variable and independent of X_1, X_2, \dots, X_N . In the real applications, X_1, X_2, \dots, X_N can be correlated, such as in insurance, finance, sports, cellular mobile systems, Internet traffic, and communications. Some authors have discussed the distributions of the sum of correlated random variables and the estimation of its parameters (for example, see Kaas et al. [4], Mehta et al. [9], Chatelain et al. [2], and Gao et al. [3]). Unfortunately, they assumed that N be fixed not a discrete random variable. Whereas, in some applications, for examples in insurance and sports, N can be a discrete random variable.

In this paper, phrase correlated random sum (CRS) will be used for the sum of correlated random variables in case where N be a discrete random variable and independent of X_1, X_2, \dots . The formulas to calculate mean and variance of CRS have been derived by Mutaqin et al. [10]. They set some assumptions on the moments of N and X . The formulas have been numerically verified by comparing between the empirical moments and theoretical moments of CRS via Monte Carlo simulation.

Mutaqin et al. [11] have proposed three approaches to estimate the mean and variance of CRS. The first approach is based on the empirical moments of CRS. The second approach is based on the empirical moments of X_j 's. The third approach is based on the weighted mean of the sum of empirical moments of X_j 's. Monte Carlo simulation is used to compare the performance of the approaches.

This paper derives the distribution, percentile, mean, and variance of CRS when X_1, X_2, \dots, X_N are correlated lognormal random variables. Therefore, we can call that Z_N as a *random sum* of correlated lognormal random variables or *correlated lognormal random sum (CLRS)*. It is assumed that conditional distribution of X_1, X_2, \dots, X_N given $N = l$ is multivariate lognormal. The multivariate lognormal distribution has widespread

application, such as in insurance, social sciences, economics, engineering, environment, and physical sciences. We use research results of Mehta et al. [9] to derive the distribution of CLRS. The Monte Carlo simulation will be conducted to validate the proposed distribution, mean, variance and percentile.

The rest of the paper is organized as follows. Section 2 presents the CLRS. The distribution and its properties of CLRS are given in Section 3. Section 4 provides the numerical results. The conclusions follow in the last section.

2. The CLRS

Let correlated random sum

$$Z_N = X_1 + X_2 + \cdots + X_N, \quad N = 0, 1, 2, \dots,$$

where $Z_0 = 0$, X_1, X_2, \dots, X_N are correlated random variables, N be a discrete random variable with the probability mass function, $p_N(n) = P(N = n)$ for $n = 0, 1, 2, \dots$, and independent of X_1, X_2, \dots . CLRS is constructed by assuming that the distribution of X_1, X_2, \dots, X_N given $N = l$ is multivariate lognormal, with probability density function

$$\begin{aligned} & f(x_1, x_2, \dots, x_l) \\ &= \frac{1}{(2\pi)^{l/2} |\mathbf{\Lambda}|^{1/2} (x_1 x_2 \cdots x_l)} \exp \left\{ -\frac{1}{2} (\ln \mathbf{x} - \boldsymbol{\theta})^T \mathbf{\Lambda}^{-1} (\ln \mathbf{x} - \boldsymbol{\theta}) \right\}, \\ & \quad x_i > 0, \quad i = 1, 2, \dots, l, \end{aligned}$$

where

$$\mathbf{x} = (x_1, x_2, \dots, x_l)^T, \quad \ln \mathbf{x} = (\ln x_1, \ln x_2, \dots, \ln x_l)^T,$$

$$\boldsymbol{\theta} = (\theta_1, \theta_2, \dots, \theta_l)^T,$$

and $\mathbf{\Lambda} = (\lambda_{i,j})$. More details about multivariate lognormal distribution can

be found in Kotz et al. [7, pp. 27-28, 219-230], and Kleiber and Kotz [5, pp. 124-126].

3. Distribution of CLRS and Its Properties

Mehta et al. [9] showed that the distribution of $Z_N | N$ can be approximated by the univariate lognormal distribution. Let μ_{Z_l} and $\sigma_{Z_l}^2$ denote the parameters of the distribution of Z_l (or $Z_N | N = l$), then those parameters can be estimated numerically using the following two equations:

$$\hat{\Psi}_{Z_l}(s_i; \mu_{Z_l}, \sigma_{Z_l}^2) = \hat{\Psi}_{\sum_{j=1}^l X_j}^{(c)}(s_i; \boldsymbol{\theta}, \mathbf{C}); \quad i = 1, 2,$$

where

$$\hat{\Psi}_{Z_l}(s_i; \mu_{Z_l}, \sigma_{Z_l}^2) \approx \sum_{k=1}^M \frac{w_k}{\sqrt{\pi}} \exp[-s_i \exp(\sqrt{2}\sigma_{Z_l}a_k + \mu_{Z_l})], \quad (1)$$

and

$$\begin{aligned} & \hat{\Psi}_{\sum_{j=1}^l X_j}^{(c)}(s_i; \boldsymbol{\theta}, \mathbf{C}) \\ & \approx \sum_{k_1=1}^M \cdots \sum_{k_l=1}^M \left[\prod_{t=1}^l \frac{w_{k_t}}{\sqrt{\pi}} \right] \exp \left[-s_i \sum_{t=1}^l \left[\exp \left(\sqrt{2} \sum_{j=1}^l c_{tj} a_{k_j} + \theta_t \right) \right] \right]. \quad (2) \end{aligned}$$

The equations (1) and (2) are the approximation moment generating function using Gauss-Hermite expansion for real s_i of Z_l and $\sum_{j=1}^l X_j$, respectively, M is the Hermite integration order. The value of $M = 12$ was found to be accurate for the correlated case (Mehta et al. [9]). The weights, w_k , and abscissas, a_k , for M up to 20 are tabulated in Abramowitz and Stegun [1, Tbl. 25.10, p. 924]. The c_{ij} is the (i, j) th element of matrix \mathbf{C} , where \mathbf{C} be the square root of the matrix $\boldsymbol{\Lambda}$, i.e., $\boldsymbol{\Lambda} = \mathbf{C}\mathbf{C}^T$.

According to the results of Mehta et al. [9], it is easy to show that the approximation distribution function of CLRS, Z_N , is given by

$$\tilde{F}(z) = \begin{cases} 0; & z < 0, \\ p_N(0); & z = 0, \\ p_N(0) + \sum_{l=1}^{\infty} p_N(l) \Phi\left(\frac{\ln z - \mu_{Z_l}}{\sigma_{Z_l}}\right); & z > 0, \end{cases} \quad (3)$$

where $\Phi(t)$ denote the standard normal distribution function evaluated at t .

The approximation probability density function of CLRS, Z_N , is given by

$$\tilde{f}(z) = \begin{cases} p_N(0); & z = 0, \\ \sum_{l=1}^{\infty} \frac{p_N(l)}{z \sigma_{Z_l} \sqrt{2\pi}} \exp\left[-\frac{1}{2} \left(\frac{\ln z - \mu_{Z_l}}{\sigma_{Z_l}}\right)^2\right]; & z > 0. \end{cases} \quad (4)$$

If N is constant, and $N > 0$, then the resulting approximation distribution of Z_N is univariate lognormal distribution. The approximation expectation of CLRS, Z_N , is given by

$$\begin{aligned} \tilde{E} &= E[Z_N] = \int_0^{\infty} z f_{Z_N}(z) dz \\ &\approx 0 p_N(0) + \int_0^{\infty} z \sum_{l=1}^{\infty} \frac{p_N(l)}{z \sigma_{Z_l} \sqrt{2\pi}} \exp\left[-\frac{1}{2} \left(\frac{\ln z - \mu_{Z_l}}{\sigma_{Z_l}}\right)^2\right] dz \\ &= \sum_{l=1}^{\infty} p_N(l) \int_0^{\infty} \frac{z}{z \sigma_{Z_l} \sqrt{2\pi}} \exp\left[-\frac{1}{2} \left(\frac{\ln z - \mu_{Z_l}}{\sigma_{Z_l}}\right)^2\right] dz \\ &= \sum_{l=1}^{\infty} p_N(l) \exp\left(\mu_{Z_l} + \frac{\sigma_{Z_l}^2}{2}\right). \end{aligned} \quad (5)$$

The approximation variance of CLRS, Z_N , is given by

$$\begin{aligned}
 \tilde{V} &= \text{Var}[Z_N] = E[Z_N^2] - (E[Z_N])^2 \\
 &= \int_0^\infty z^2 \sum_{l=1}^\infty \frac{p_N(l)}{z\sigma_{Z_l}\sqrt{2\pi}} \exp\left[-\frac{1}{2}\left(\frac{\log z - \mu_{Z_l}}{\sigma_{Z_l}}\right)^2\right] dz - \tilde{E}^2 \\
 &= \sum_{l=1}^\infty p_N(l) \int_0^\infty \frac{z^2}{z\sigma_{Z_l}\sqrt{2\pi}} \exp\left[-\frac{1}{2}\left(\frac{\log z - \mu_{Z_l}}{\sigma_{Z_l}}\right)^2\right] dz - \tilde{E}^2 \\
 &= \sum_{l=1}^\infty p_N(l) \exp(2\mu_{Z_l} + 2\sigma_{Z_l}^2) - \tilde{E}^2.
 \end{aligned} \tag{6}$$

Let \tilde{z}_p denotes the approximation 100pth percentile of Z_N . Then

$$F_{Z_N}(\tilde{z}_p) = \Pr(Z_N \leq \tilde{z}_p) = p.$$

Therefore $\tilde{z}_p = 0$ for $p \leq p_N(0)$, or \tilde{z}_p is the unique solution of the following equation:

$$\sum_{l=1}^\infty p_N(l) \Phi\left(\frac{\ln \tilde{z}_p - \mu_{Z_l}}{\sigma_{Z_l}}\right) = p - p_N(0), \tag{7}$$

for $p > p_N(0)$.

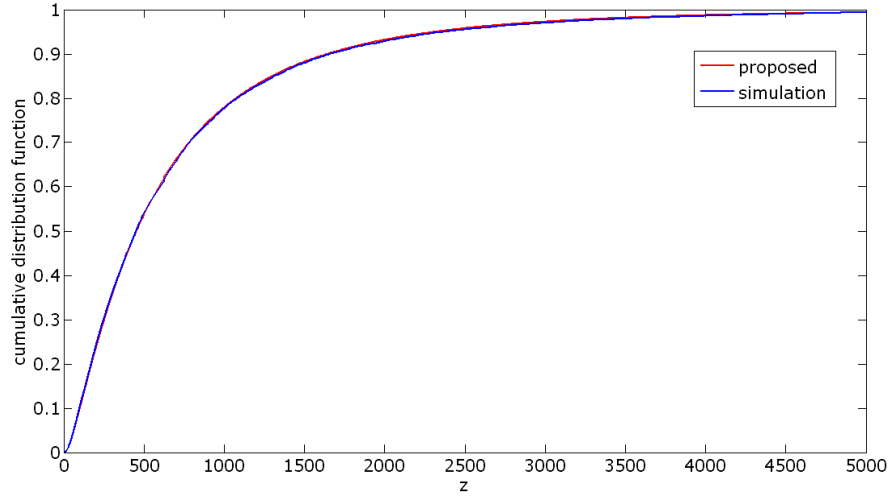


Figure 1. Comparison between the proposed distribution and simulation for $N \sim DU(1, 5)$.

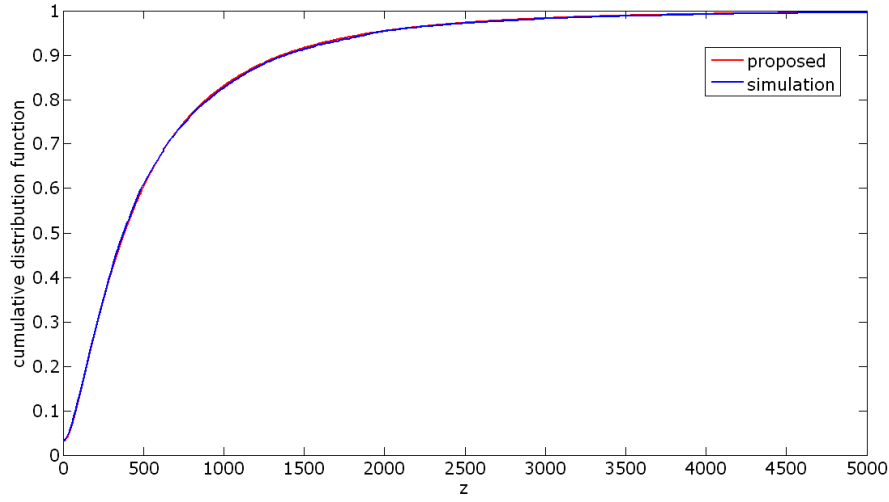


Figure 2. Comparison between the proposed distribution and simulation for $N \sim bin(3, 0.5)$.

4. Numerical Results

This section contains numerical results to validate the distribution of CLRS and its properties are discussed in Section 3. Monte Carlo simulation

is conducted 10.000 times to simulate CLRS, where the distribution of X_1, X_2, \dots, X_N given $N = l$ is multivariate lognormal, with parameters $\theta_i = 5$, $\lambda_{i,i} = 1$, and $\lambda_{i,j} = 0$, 62, for $i \neq j = 1, 2, \dots$. Empirical distribution from simulated CLRS can give an appropriate approximation to exact distribution of CLRS. Figure 1 and Figure 2 show comparison between the proposed approximation cumulative distribution function and empirical distribution from simulated CLRS, where the distributions of N are discrete uniform, $DU(1, 5)$ (Law and Kelton [8]), and binomial, $bin(5, 0.5)$ (Law and Kelton [8]), respectively. Comparisons between the proposed approximation 100 p th percentile and simulation are displayed in Table 1 and Table 2. Based on Figures 1-2, and Tables 1-2, we can conclude that the proposed distribution of CLRS is not too different from the exact distribution of CLRS.

Table 1. Comparison between the proposed approximation 100 p th percentile and simulation for distribution of N is discrete uniform

p	$N \sim DU(1, 3)$		$N \sim DU(1, 5)$	
	Proposed	Simulation	Proposed	Simulation
0.10	72.62	72.80	100.98	98.25
0.25	144.09	143.00	211.29	207.01
0.50	299.03	293.33	451.28	449.00
0.75	598.01	596.52	907.38	917.67
0.90	1087.86	1102.73	1642.22	1656.67

Table 2. Comparison between the proposed approximation 100 p th percentile and simulation for distribution of N is binomial

p	$N \sim bin(3; 0.5)$		$N \sim bin(5; 0.5)$	
	Proposed	Simulation	Proposed	Simulation
0.10	0.00	0.00	80.84	78.71
0.25	79.64	80.81	180.14	179.19
0.50	208.74	208.06	380.20	374.91
0.75	454.55	458.10	755.43	763.66
0.90	863.85	875.79	1363.09	1384.38

Tables 3-4 display comparison between the proposed approximation values and its exact values for mean and variance of CLRS. It can be seen that the proposed approximation values are not too different from its exact values. These proofs show that the formulas of the proposed approximation mean and variance of CLRS are valid.

Table 3. Comparison between the proposed approximation and simulation for mean and variance of CLRS when the distribution of N is discrete uniform

Moment	$N \sim DU(1, 3)$		$N \sim DU(1, 5)$	
	Exact	Approximation	Exact	Approximation
Mean	8.9634	8.9618	13.4451	13.4427
Variance	128.4210	127.8852	281.7252	280.7758

Table 4. Comparison between the proposed approximation and simulation for mean and variance of CLRS when the distribution of N is binomial

Moment	$N \sim bin(3; 0.5)$		$N \sim bin(5; 0.5)$	
	Exact	Approximation	Exact	Approximation
Mean	6.7225	6.7216	11.2042	11.2021
Variance	92.7111	92.3718	197.6486	196.8718

5. Conclusions

We have proposed a simple method to approximate the distribution of CLRS. The resulting distribution is mixture of lognormal distribution, when $N > 0$. The result of Mehta et al. [9] is special case of the resulting distribution when N is constant. We also have derived percentile, mean and variance of CLRS. The numerical results showed that the distribution, percentile, mean and variance of CLRS are valid and good approximations.

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